

# COMMONWEALTH OF AUSTRALIA

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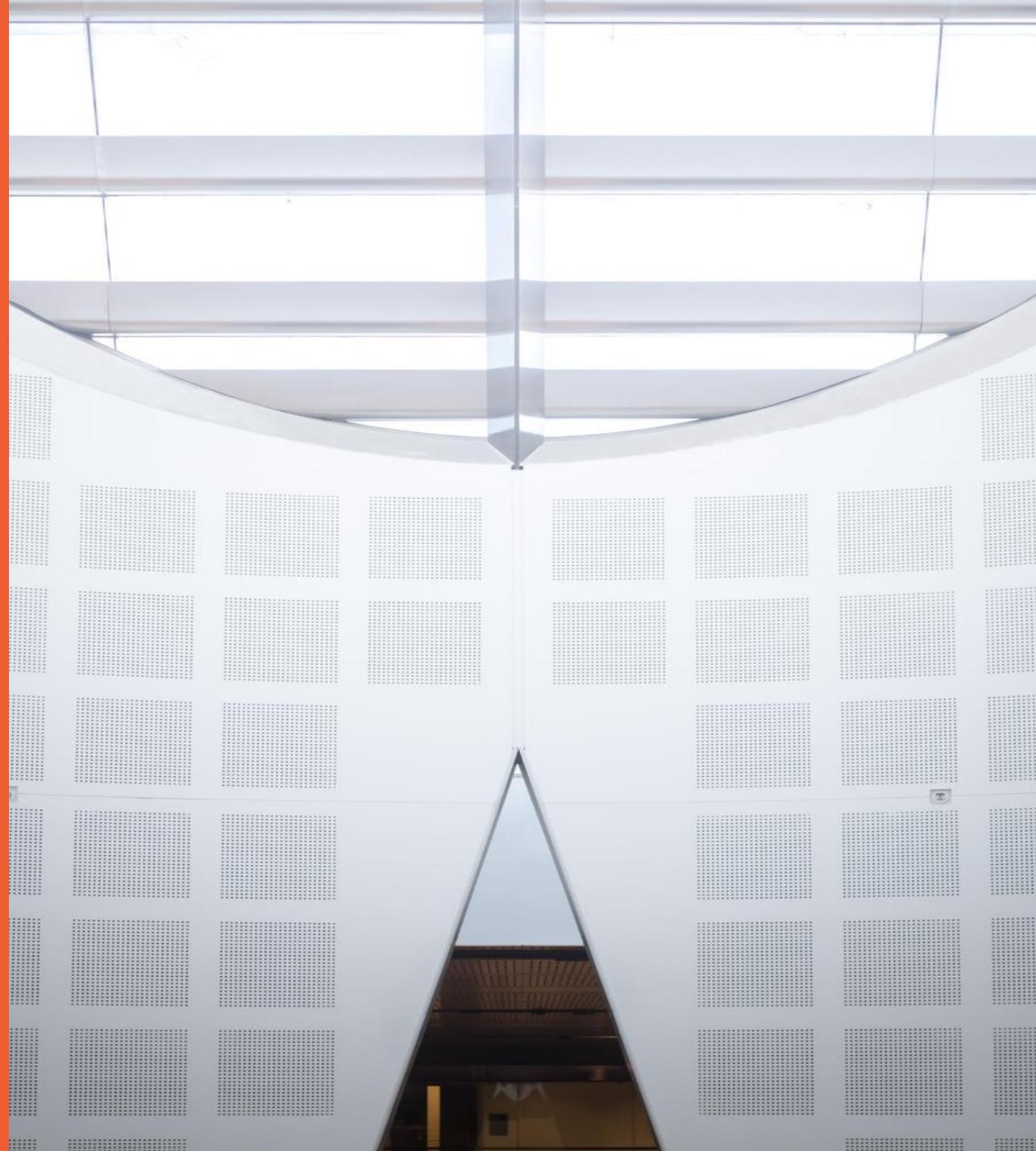
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COMPx270: Randomised and  
Advanced Algorithms  
Lecture 11: Learning and testing  
probability distributions

Clément Canonne  
School of Computer Science



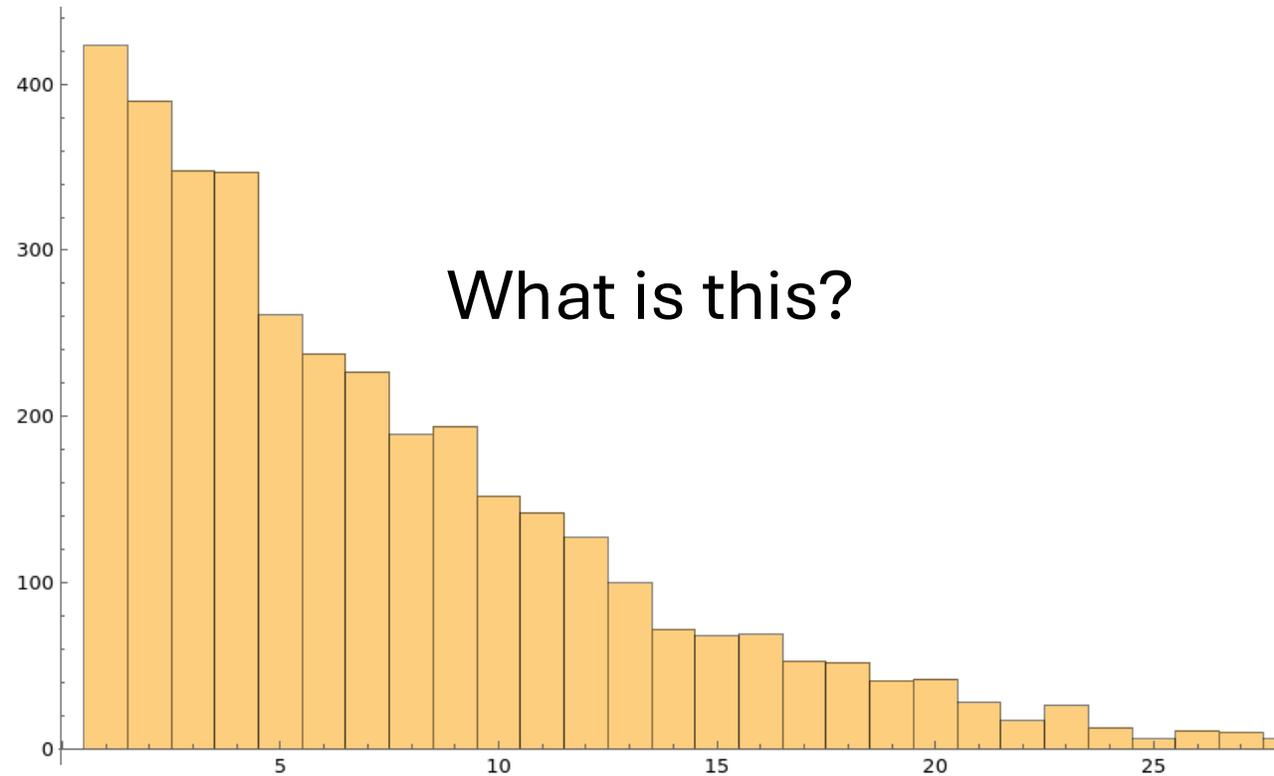
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## Some housekeeping

- A2 **still** being marked: deepest apologies (my fault)
- A3 (after Simple Extension) due **tomorrow**
- **Don't forget the "participation" assignment (Oct 18)**
- **Sample exam** is out, will be the topic of Week 13
- **Feedback** welcome: <https://forms.office.com/r/DymMcfN47n>
- **Final exam on Tues, Nov 12 (9am)** → *what is allowed?*

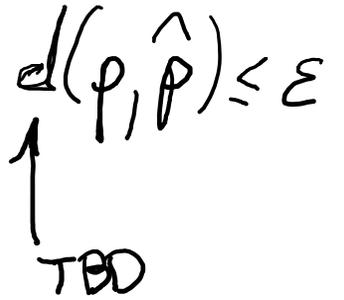
# A question



# Learning and testing (discrete) probability distributions <sup>\*</sup>

$p$  over  $\mathcal{X}$  (discrete, finite domain) of size  $k$

- learn  $p \rightarrow$  get  $\hat{p}$  (distr over  $\mathcal{X}$ ) (k numbers) st  $d(p, \hat{p}) \leq \epsilon$
- learn a parameter of  $p$  (estimation) : a number
- learn a bit (test if  $p$  satisfies some property)



"TBD" :  $d(p, q) =$  total variation distance  
(quantifies how far  $p, q$  are)

Access  $x_1, \dots, x_n \stackrel{iid}{\sim} p$  (algo only sees  $x_1, \dots, x_n$ )

\*  $p: \mathcal{X} \rightarrow [0, 1]$

$$\sum_{i \in \mathcal{X}} p(i) = 1$$

# Preliminaries on probability distributions

TV distance

$p, q$  over  $\mathcal{X}$

A metric would be nice

Bounded (in  $[0, 1]$ ?) also?

Meaning?

$$\text{TV}(p, q) = \sup_{S \subseteq \mathcal{X}} (p(S) - q(S)) = \sup_{S \subseteq \mathcal{X}} |p(S) - q(S)|$$

$$p(S) = \sum_{i \in S} p(i)$$

$$= \sup_{\substack{\text{any algo } x \sim p \\ A: \mathcal{X} \rightarrow \{0, 1\}}} (\Pr[A(x)=1] - \Pr[A(x)=1]_{x \sim q})$$

## Preliminaries on probability distributions

Fact

$$TV(p, q) = \frac{1}{2} \sum_{i \in \mathcal{X}} |p(i) - q(i)| = \frac{1}{2} \|p - q\|_1$$

Pf

$$\text{Take } S^* = \{x : p(x) > q(x)\}$$

$$TV(p, q) \geq p(S^*) - q(S^*) = \sum_{i \in S^*} (p(i) - q(i)) = \sum_{i \in S^*} |p(i) - q(i)|$$

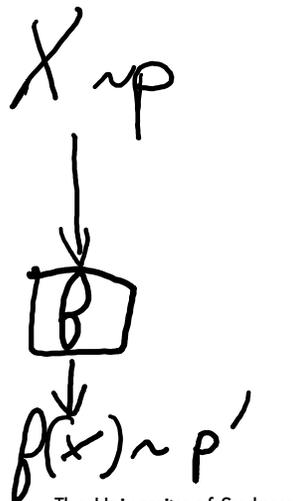
$$\begin{aligned} \text{But } \sum_{i \in \mathcal{X}} |p(i) - q(i)| &= \sum_{i \in S^*} (p(i) - q(i)) + \sum_{i \notin S^*} (q(i) - p(i)) \\ &= 2 \sum_{i \in S^*} (p(i) - q(i)) \\ &= 2 (p(S^*) - q(S^*)) \end{aligned}$$
$$\begin{aligned} &\underbrace{\sum_{i \notin S^*} q(i) - \sum_{i \notin S^*} p(i)} \\ &= 1 - q(S^*) - (1 - p(S^*)) \end{aligned}$$

# Preliminaries on probability distributions

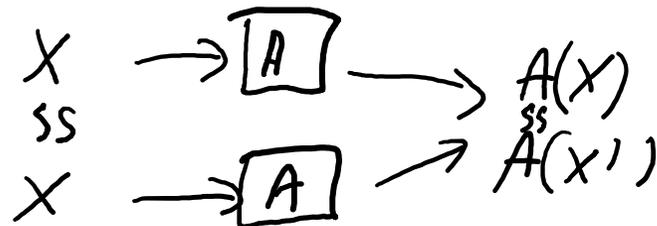
DPI (Data Processing Inequality)

Take any  $f: X \rightarrow Y$

If  $X \sim p$  let  $p'$  be the distr of  $f(X)$   
 $X' \sim q$  let  $q'$  —————  $f(X')$



$$TV(p', q') \leq TV(p, q)$$



## A view of TV distance

Alice and Bob play a game, where they both know two probability distributions  $\mathbf{p}, \mathbf{q}$ . Alice starts by tossing a fair coin, and does not show the outcome to Bob: if it is Heads, then she draws  $x \sim \mathbf{p}$ ; if it is Tails, she draws  $x \sim \mathbf{q}$ . Then she shows the value of  $x$  to Bob, who must guess if the coin toss was Heads. Clearly, just by random guessing, Bob can win the game with probability  $1/2$ . What the lemma says is that he can do better: there is a strategy for him to win with probability

$$\Pr[\text{Bob wins}] = \frac{1}{2} + \frac{d_{\text{TV}}(\mathbf{p}, \mathbf{q})}{2}$$

and, moreover, this is the best possible.

## The case of a coin $\square$

How many times  $n$  do you need to flip the coin to learn its true bias  $p$  to accuracy  $\pm\epsilon$ , and be correct with probability at least  $1 - \delta$ ?

## The case of a coin $\square$

**Theorem 50.** *Suppose we are promised that the true bias  $p$  of the coin satisfies  $0 \leq p < q \leq \frac{1}{2}$ , for some known value  $q$ . Then estimating the bias of the coin to an additive  $\varepsilon$ , with probability at least  $1 - \delta$ , can be done with  $n = O\left(\frac{q}{\varepsilon^2} \log \frac{1}{\delta}\right)$  i.i.d. samples. (Moreover, this is optimal.)*

## The case of a coin $\square$

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**Corollary 50.1.** *Estimating the bias of a coin to an additive  $\varepsilon$ , with probability at least  $1 - \delta$ , can be done with  $n = O\left(\frac{1}{\varepsilon^2} \log \frac{1}{\delta}\right)$  i.i.d. samples. (Moreover, this is optimal.)*

# The case of a coin $\square$

**Theorem 50.** Suppose we are promised that the true bias  $p$  of the coin satisfies  $0 \leq p < q \leq \frac{1}{2}$ , for some known value  $q$ . Then estimating the bias of the coin to an additive  $\epsilon$ , with probability at least  $1 - \delta$ , can be done with  $n = O\left(\frac{q}{\epsilon^2} \log \frac{1}{\delta}\right)$  i.i.d. samples. (Moreover, this is optimal.)

$$\hat{p} = \frac{1}{n} \sum_{i=1}^n x_i$$

Hoeffding  $\downarrow$  want

$$\Pr[|\hat{p} - p| > \epsilon] \leq 2e^{-2\epsilon^2 n} \leq \delta$$

take  $n = \left\lceil \frac{1}{2\epsilon^2} \ln \frac{2}{\delta} \right\rceil$

Chernoff (if  $q > \epsilon$  (if not, easy...))  $\uparrow$  want

$$\leq 2e^{-\frac{\epsilon^2 q n}{3}} \leq \delta$$

$$n = \left\lceil \frac{30}{\epsilon^2 q} \ln \frac{2}{\delta} \right\rceil$$

## The case of a coin: what about testing?

Is my coin a fair coin?

# The case of a coin: what about testing? $\square$

**Theorem 51.** Testing whether the bias of a coin is  $1/2$  or at least  $1/2 + \varepsilon$ , with probability at least  $1 - \delta$ , can be done with  $n = O\left(\frac{1}{\varepsilon^2} \log \frac{1}{\delta}\right)$  i.i.d. samples. (Moreover, this is optimal.)

Not better than learning  
P.

# The case of a coin: what about testing? $\square$

**Theorem 52.** For any  $0 < \alpha \leq 1/2$  and  $\varepsilon \in (0, 1]$ , testing whether the bias of a coin is at most  $\alpha$  or at least  $\alpha(1 + \varepsilon)$ , with probability at least  $1 - \delta$ , can be done with  $n = O\left(\frac{1}{\alpha\varepsilon^2} \log \frac{1}{\delta}\right)$  i.i.d. samples.

$\alpha = \frac{1}{2} \rightarrow$  previous theorem

$\alpha$  small,  
 $\varepsilon = 1$

$\leq \alpha$   
vs  $\geq 2\alpha$ :  $O\left(\frac{1}{\alpha} \log \frac{1}{\delta}\right)$

# The case of a coin: what about testing? $\square$

**Theorem 52.** For any  $0 < \alpha \leq 1/2$  and  $\epsilon \in (0, 1]$ , testing whether the bias of a coin is at most  $\alpha$  or at least  $\alpha(1 + \epsilon)$ , with probability at least  $1 - \delta$ , can be done with  $n = O\left(\frac{1}{\alpha\epsilon^2} \log \frac{1}{\delta}\right)$  i.i.d. samples.

Can prove via Chernoff:

Compute

threshold

yes  $\alpha$   $\alpha(1 + \epsilon)$  no

$\hat{p} = \frac{1}{n} \sum_{i=1}^n x_i$

Output "yes" if  $p \leq \alpha(1 + \frac{\epsilon}{2})$

## Beyond coins: $k$ is large

Domain sizes grow quite fast, and in most settings  $k$  is huge.

## Learning in TV distance

Unknown  $p$  over  $\mathcal{X}$  ( $|\mathcal{X}| = k$ )

Parameter  $\epsilon, \delta$

Get  $x_1, \dots, x_n \stackrel{iid}{\sim} p$  for  $n$  to be chosen

Goal: output  $\hat{p}$  (distrib. over  $\mathcal{X}$ )  
such that

$$\Pr[TV(p, \hat{p}) > \epsilon] \leq \delta$$

Minimize  $\underline{n}$  (sample complexity)

$$\hookrightarrow n(k, \epsilon, \delta)$$

# Learning in TV distance: first attempt

$$P = (p_1, p_2, \dots, p_k)$$

Want:  $\hat{p} \in [0, 1]^k$  st

$$\star \|\hat{p} - p\|_1 \leq 2\varepsilon \quad (\text{w/p } \geq 1 - \delta)$$

(If  $\|\hat{p}\|_1 \neq 1$ ,  
can normalise)

$$\sum_{i=1}^k |\hat{p}_i - p_i|$$

$\star$  implies

$$|\|\hat{p}\|_1 - \underbrace{\|p\|_1}_{=1}| \leq 2\varepsilon$$

Enough to estimate each  $p_i$  by  $\hat{p}_i$  to  $\pm \frac{2\varepsilon}{k}$   
w/p  $1 - \frac{\delta}{k}$  each

$k$  instances of "estimate bias  $p_i$  of a coin"

Gives us

$$n = O\left(\frac{1}{(\varepsilon/k)^2} \log \frac{1}{(\delta/k)}\right) = O\left(\frac{k^2}{\varepsilon^2} \log \frac{1}{\delta}\right)$$

## Learning in TV distance: second attempt

What if instead of  $\hat{p}_i \approx p_i \pm \frac{2\varepsilon}{k}$ , we try  $\hat{p}_i \approx (1 \pm 2\varepsilon) p_i$ .

Then 
$$\|\hat{p} - p\|_1 = \sum_{i=1}^k |p_i - \hat{p}_i| \leq \sum_{i=1}^k 2\varepsilon p_i = 2\varepsilon$$

\* If  $\min_{1 \leq i \leq k} p_i \geq \frac{\varepsilon}{k}$ , we can (if no lower bound on  $p_i$ , can't)

Chernoff bound + union bound ( $\delta' = \frac{\delta}{k}$ )

$$\left( \leq e^{-\varepsilon^2 \frac{k}{\delta}} \right)^n$$
  
when using Chernoff

$$n = O\left(\frac{k}{\varepsilon^3} \log \frac{k}{\delta}\right)$$

\* replace  $p$  by  $p' = (1 - \frac{\varepsilon}{2})p + \frac{\varepsilon}{2}u_R$   
Learn  $p'$  to TV  $\frac{\varepsilon}{2}$ .

# Learning in TV distance: third attempt

**Theorem 53.** Learning an unknown distribution  $\mathbf{p} \in \Delta(k)$  to total variation distance  $\varepsilon$  (with success probability  $1 - \delta$ ) can be done with

$$n = O\left(\frac{k + \log \frac{1}{\delta}}{\varepsilon^2}\right)$$

*i.i.d. samples. (Moreover, this is optimal.)*

↓  
The empirical estimator works  
 $\hat{p}_i = \frac{N_i}{n} \leftarrow \# \text{ of times } i \text{ is seen in the } n \text{ samples}$

$$n = O\left(\frac{k + \log \frac{1}{\delta}}{\epsilon^2}\right)$$

i.i.d. samples. (Moreover, this is optimal.)

# Learning in TV distance: third attempt

Want  $TV(p, \hat{p}) \leq \epsilon$

$$TV(p, \hat{p}) > \epsilon \Leftrightarrow \exists S \subseteq \mathcal{X}, |p(S) - \hat{p}(S)| > \epsilon$$

For any  $S \subseteq \mathcal{X}$ .

$$Pr[|p(S) - \hat{p}(S)| > \epsilon] \leq 2e^{-2\epsilon^2 n} \stackrel{\text{want}}{\leq} \frac{\delta}{2^k}$$

this is Hoeffding.  
Bias of a coin!

$$p(S) = \sum_{i \in S} p_i$$

$$\hat{p}(S) = \sum_{i \in S} \hat{p}_i$$

$$= \frac{\text{\# samples falling in } S}{n}$$

$$n = \left\lceil \frac{\ln\left(\frac{2^{k+1}}{\delta}\right)}{2\epsilon^2} \right\rceil = O\left(\frac{k + \log\left(\frac{1}{\delta}\right)}{\epsilon^2}\right)$$

+ union bound over  $2^k$  subsets.

Theorem 53. Learning an unknown distribution  $\mathbf{p} \in \Delta(k)$  to total variation distance  $\epsilon$  (with success probability  $1 - \delta$ ) can be done with

$$n = O\left(\frac{k + \log \frac{1}{\delta}}{\epsilon^2}\right)$$

i.i.d. samples. (Moreover, this is optimal.)

## Learning in TV distance: second third attempt

$$\begin{aligned} \mathbb{E}[\text{TV}(\mathbf{p}, \hat{\mathbf{p}})] &= \frac{1}{2} \mathbb{E}[\|\mathbf{p} - \hat{\mathbf{p}}\|_1] = \frac{1}{2} \sum_{i=1}^k \mathbb{E}[|p_i - \hat{p}_i|] \\ &\stackrel{\text{Jensen}}{\leq} \frac{1}{2} \sum_{i=1}^k \sqrt{\mathbb{E}[(p_i - \hat{p}_i)^2]} = \frac{1}{2n} \sum_{i=1}^k \sqrt{\text{Var}(n\hat{p}_i)} \\ &\quad \text{Var } \hat{p}_i \quad \underbrace{\text{Var}(n\hat{p}_i)}_{np_i(1-p_i) \leq np_i} \end{aligned}$$

Jensen  
 $f(x) = x^2$

$$\mathbb{E}\hat{p}_i = p_i$$

$$n\hat{p}_i \sim \text{Bin}(n, p_i)$$

$$\leq \frac{1}{2\sqrt{n}} \sum_{i=1}^k \sqrt{p_i} \stackrel{\text{Jensen}}{\leq} \frac{\sqrt{k}}{2\sqrt{n}}$$

(Jensen)

want  $\leq \frac{\epsilon}{10}$

true for  $n \geq \frac{25k}{\epsilon^2}$   
+ Markov.

$$\begin{aligned} k \cdot \frac{1}{k} \sum_{i=1}^k \sqrt{p_i} &= k \mathbb{E} \sqrt{p_i} \\ &\leq k \sqrt{\mathbb{E} p_i} \\ &= k \sqrt{\frac{1}{k} \sum_{i=1}^k p_i} = \sqrt{k} \end{aligned}$$

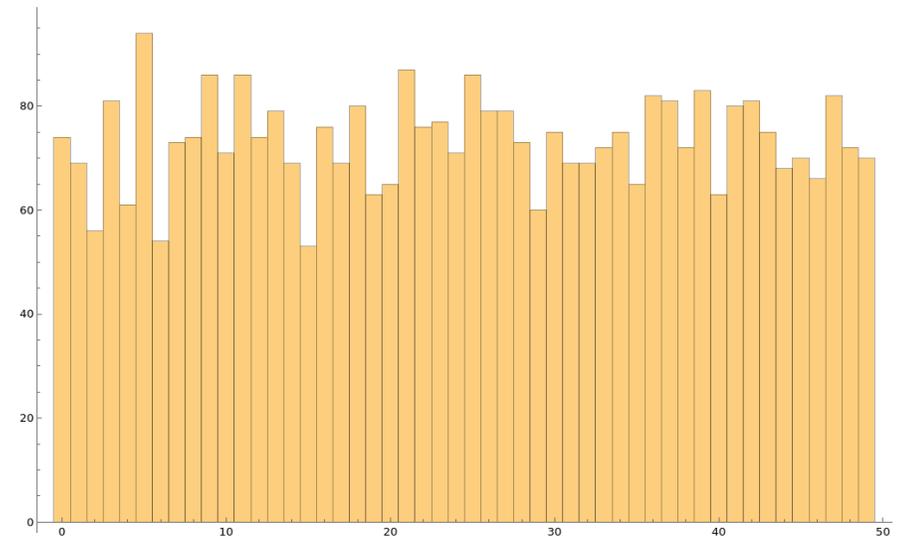
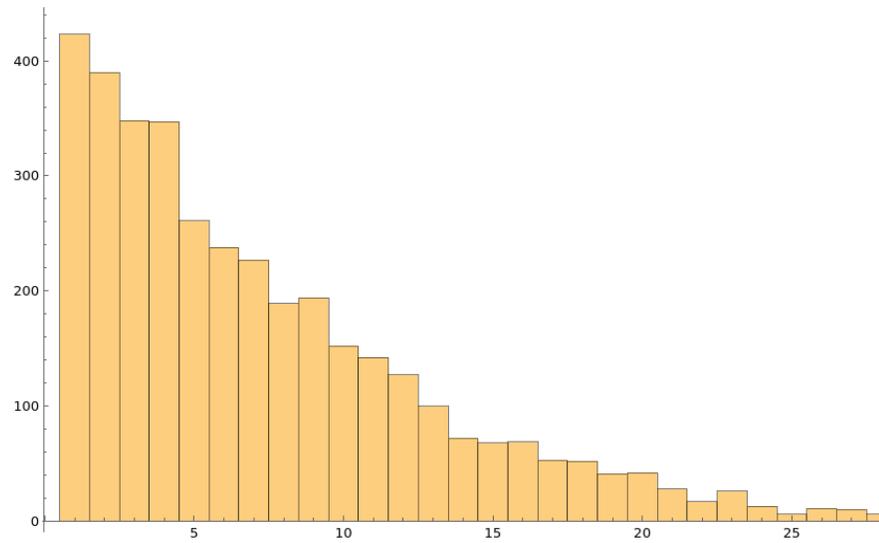
# Learning in TV distance: second third attempt

**Theorem 53.** Learning an unknown distribution  $\mathbf{p} \in \Delta(k)$  to total variation distance  $\varepsilon$  (with success probability  $1 - \delta$ ) can be done with

$$n = O\left(\frac{k + \log \frac{1}{\delta}}{\varepsilon^2}\right)$$

*i.i.d. samples. (Moreover, this is optimal.)*

# Testing in TV distance



# Testing in TV distance: identity testing

Give an algorithm  $A$  which takes parameters  $\epsilon, \delta \in (0, 1]$  and  $n$  samples from  $\mathbf{p}$ , and:

- If  $\mathbf{p} = \mathbf{q}$ , then  $\Pr[A \text{ outputs yes}] \geq 1 - \delta$ ;
- If  $d_{\text{TV}}(\mathbf{p}, \mathbf{q}) > \epsilon$ , then  $\Pr[A \text{ outputs no}] \geq 1 - \delta$

(if  $0 < d_{\text{TV}}(\mathbf{p}, \mathbf{q}) \leq \epsilon$ , then  $A$  is off the hook and can output whatever).



## Testing in TV distance: identity testing via learning

$$n = O\left(\frac{k + \log(1/\delta)}{\epsilon^2}\right) \text{ is an upper bound}$$

$p \rightarrow \hat{p}$  st  $TV(p, \hat{p}) \leq \frac{\epsilon}{2} \rightarrow$  check if  $TV(\hat{p}, q) \leq \frac{\epsilon}{2} \rightarrow$  if not, say "no"

$$\text{Key: } TV(p, q) \leq TV(p, \hat{p}) + TV(\hat{p}, q)$$

# Testing in TV distance: uniformity is all you need

**Theorem 54** (Identity to uniformity reduction). Suppose there is an algorithm  $A$  for uniformity testing, which takes  $n = n(k, \epsilon, \delta)$  i.i.d. samples from the unknown distribution. Then there is an algorithm  $A'$  for identity testing over a domain of size  $k$  to any fixed  $\mathbf{q} \in \Delta(k)$ , which takes  $n = n(4k, \epsilon/4, \delta)$  i.i.d. samples from the unknown distribution. Moreover,  $A'$  is efficient if  $A$  is.

Upshot: "wlog,  $p = u_k$   
vs  
 $TV(p, u_k) > \epsilon$ "  
(no need to worry about other  $q$ )

# Testing in TV distance: uniformity testing

Uniform is simpler

(Focus on  $\mathcal{D} = \frac{1}{3}$ )

Birthday paradox:

If  $p$  is uniform on  $\frac{k}{2}$  elem<sup>ts</sup> }  $n = \Omega(\sqrt{k})$   
vs  $p$  is the uniform distr (over  $k$  elem<sup>ts</sup>) }

# Testing in TV distance: uniformity testing

**Theorem 55.** *Testing uniformity of an unknown distribution  $\mathbf{p} \in \Delta(k)$  to total variation distance  $\varepsilon$  (with success probability  $2/3$ ) can be done with*

$$n = O\left(\frac{\sqrt{k}}{\varepsilon^2}\right)$$

*i.i.d. samples, using Algorithm 21. (Moreover, this is optimal for constant success probability.)*

## Testing in TV distance: uniformity testing, key ideas

$$\text{TV} \rightarrow l_2$$

$$\|p - u_h\|_1 \leq \sqrt{k} \|p - u_h\|_2$$

$$\text{so if } p = u_h, \quad \|p - u_h\|_2 = 0$$

$$\text{if } \text{TV}(p, u_h) > \varepsilon, \quad \|p - u_h\|_2 > \frac{2\varepsilon}{\sqrt{k}}$$

## Testing in TV distance: uniformity testing, key ideas

$$\textcircled{2} \quad \left\| p - u_k \right\|_2^2 = \left\| p \right\|_2^2 - \frac{1}{k}$$

$$\sum_i \left( p_i - \frac{1}{k} \right)^2 = \sum_i p_i^2 - \frac{2}{k} \left( \sum_i p_i \right) + \frac{1}{k} = \left\| p \right\|_2^2 - \frac{1}{k}$$

$= 1$

- $\left\| p \right\|_2^2 = \frac{1}{k}$  if  $p = u_k$

- $\left\| p \right\|_2^2 > \frac{1 + 4\varepsilon^2}{k}$  if  $\text{TV}(p, u_k) > \varepsilon$

$$\left\| p \right\|_2^2 = \Pr_{x, y \sim p} [x = y]$$

$$\sum_i \Pr[x = y = i] = \sum_i p_i \cdot p_i$$

# Testing in TV distance: uniformity testing, algorithm

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**Input:** Multiset of  $n$  i.i.d. samples  $x_1, \dots, x_n \in \mathcal{X}$ , parameters  $\varepsilon \in (0, 1]$  and  $k = |\mathcal{X}|$

1: Set  $\tau \leftarrow \frac{1+2\varepsilon^2}{k}$

2: Compute

▷  $O(n)$  time if  $\mathcal{X}$  is known

$$Z = \frac{1}{\binom{n}{2}} \sum_{1 \leq s < t \leq n} \mathbb{1}_{\{x_s = x_t\}} = \frac{1}{\binom{n}{2}} \sum_{j \in \mathcal{X}} \binom{N_j}{2}$$

where  $N_j \leftarrow \sum_{t=1}^n \mathbb{1}_{\{x_t = j\}}$ .

3: **if**  $Z \geq \tau$  **then return** no

▷ Not uniform

4: **else return** yes

▷ Uniform

---

# Testing in TV distance: uniformity testing

**Input:** Multiset of  $n$  i.i.d. samples  $x_1, \dots, x_n \in \mathcal{X}$ , parameters  $\epsilon \in$

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3: if  $Z \geq \tau$  then return no

4: else return yes

$\triangleright$  Not uniform

$\triangleright$  Uniform

①  $\mathbb{E} Z = \|p\|_2^2$  (easy)

②  $\text{Var} Z < \text{small}$  (really hard) } + Chebyshev

"Easy": can prove a bound on  $\text{Var} Z$

giving

$$n = O\left(\frac{k}{\epsilon^4}\right)$$

Getting  $O\left(\frac{\sqrt{k}}{\epsilon^2}\right)$  is hard

# Testing in TV distance: uniformity testing

---

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where  $N_j \leftarrow \sum_{i=1}^n \mathbb{1}_{\{x_i=j\}}$ .

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1: Set  $\tau \leftarrow \frac{1+2\epsilon^2}{k}$

2: Compute ▷  $O(n)$  time if  $\mathcal{X}$  is known

$$Z = \frac{1}{\binom{n}{2}} \sum_{1 \leq s < t \leq n} \mathbb{1}_{\{x_s = x_t\}} = \frac{1}{\binom{n}{2}} \sum_{j \in \mathcal{X}} \binom{N_j}{2}$$

where  $N_j \leftarrow \sum_{i=1}^n \mathbb{1}_{\{x_i=j\}}$ .

3: **if**  $Z \geq \tau$  **then return no** ▷ Not uniform

4: **else return yes** ▷ Uniform

---

# Testing in TV distance: uniformity testing, summary

**Theorem 55.** Testing uniformity of an unknown distribution  $\mathbf{p} \in \Delta(k)$  to total variation distance  $\varepsilon$  (with success probability  $2/3$ ) can be done with

$$n = O\left(\frac{\sqrt{k}}{\varepsilon^2}\right)$$

*i.i.d. samples, using Algorithm 21. (Moreover, this is optimal for constant success probability.)*

Tight bound (other algo)

$$n = O\left(\frac{\sqrt{k \log 1/\delta} + \log 1/\delta}{\varepsilon^2}\right)$$

# Summary

- Learning  $p$   
VI  $\boxed{k/\epsilon^2}$
- Testing  $p = u_k$   
vs  $TV(p, u_k) \geq \epsilon$   $\boxed{\sqrt{k}/\epsilon^2}$
- Estimating TV distance  
 $TV(p, u_k)$   $\boxed{\frac{k}{\epsilon^2 \log k}}$